

# THE LIMITING DISTRIBUTION OF THE COEFFICIENTS OF THE $q$ -CATALAN NUMBERS

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ABSTRACT. We show that the limiting distributions of the coefficients of the  $q$ -Catalan numbers and the generalized  $q$ -Catalan numbers are normal. Despite the fact that these coefficients are not unimodal for small  $n$ , we conjecture that for sufficiently large  $n$ , the coefficients are unimodal and even log-concave except for a few terms of the head and tail.

## 1. INTRODUCTION

The main objective of this paper is to show that the limiting distribution of the coefficients of the  $q$ -Catalan numbers is normal. The Catalan numbers

$$C_n = \frac{1}{n+1} \binom{2n}{n}$$

have many combinatorial interpretations, see Stanley [10]. The usual  $q$ -analog of the Catalan numbers is given by

$$(1.1) \quad C_n(q) = \frac{1}{[n+1]} \begin{bmatrix} 2n \\ n \end{bmatrix},$$

where  $[n] = 1 + q + q^2 + \cdots + q^{n-1}$ , and

$$\begin{bmatrix} n \\ k \end{bmatrix} = \frac{[n]!}{[k]![n-k]}.$$

There are also other types of  $q$ -analogs of the Catalan numbers, see, for example, Andrews [2], Gessel and Stanton [4], Krattenthaler [5].

We further consider the limiting distribution of the coefficients of the quotient of two products, which includes the result for the  $q$ -Catalan numbers as a special case. We conclude this paper with two conjectures on the unimodality and log-concavity for almost all the coefficients of the  $q$ -Catalan numbers and the generalized  $q$ -Catalan numbers provided that  $n$  is sufficiently large.

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## 2. THE LIMITING DISTRIBUTION

In this section, we use the moment generating function technique to obtain the limiting distribution of the coefficients of the  $q$ -Catalan numbers. We introduce the random variable  $\xi_n$  corresponding to the probability generating function

$$\phi_n(q) = C_n(q)/C_n.$$

As far as the computations are concerned, we will not need the following combinatorial interpretation of  $C_n(q)$ . However, for the sake of completeness, we recall that  $\xi_n$  reflects the distribution of the major indices of Catalan words of length  $2n$ , see, for example, [3]. Moreover, we write

$$C_n(q) = \sum m_n(k)q^k,$$

where  $m_n(k)$  stands for the number of Catalan words of length  $2n$  with major index  $k$ . The following lemma gives the expectation and variance of  $\xi_n$ .

**Lemma 2.1.** *We have*

$$(2.1) \quad E(\xi_n) = \frac{n(n-1)}{2} \quad \text{and} \quad \text{Var}(\xi_n) = \frac{n(n-1)(n+1)}{6}.$$

*Proof.* By the definition of  $C_n(q)$ , it is easy to check the following symmetry property of  $m_n(k)$ :

$$m_n(k) = m_n(n(n-1) - k).$$

Hence

$$E(\xi_n) = \frac{n(n-1)}{2}.$$

Let

$$F = F(q) = \prod_{i=1}^{n-1} (1 + q + \cdots + q^{n+i}) \quad \text{and} \quad G = G(q) = \prod_{i=1}^{n-1} (1 + q + \cdots + q^i).$$

It is easily verified that  $C_n(q) = F/G$ . Since

$$\begin{aligned} C_n(q)''|_{q=1} &= \left( \frac{F''}{G} - \frac{FG''}{G^2} - \frac{2G'F'}{G^2} + \frac{2G'^2F}{G^3} \right) \Big|_{q=1} \\ &= \frac{1}{12}n(n-1)(3n^2 - n - 4)C_n, \end{aligned}$$

we obtain

$$\text{Var}(\xi_n) = \frac{C_n(q)''|_{q=1}}{C_n} + E(\xi_n) - E(\xi_n)^2 = \frac{1}{6}n(n-1)(n+1).$$

This completes the proof. ■

**Lemma 2.2.** *When  $n \rightarrow \infty$ , we have*

$$\sum_{k=2}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)! \sigma^{2k}} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}) \rightarrow 0$$

uniformly for  $t$  from any bounded set, where  $B_j$ 's are the Bernoulli numbers and  $\sigma^2$  is the variance of  $\xi_n$  as given in (2.1).

*Proof.* The second summation can be expanded as follows:

$$\sum_{i=2}^n ((n+i)^{2k} - i^{2k}) = \sum_{i=2}^n \sum_{j=1}^{2k} \binom{2k}{j} n^j i^{2k-j} = \sum_{j=1}^{2k} \binom{2k}{j} \left( \sum_{i=2}^n n^j i^{2k-j} \right).$$

For  $k > 1$ , the second factor in the preceding summation is bounded by the following integral:

$$\sum_{i=2}^n n^j i^{2k-j} < n^j \int_1^{n+1} t^{2k-j} dt = n^j \cdot \frac{(n+1)^{2k-j+1} - 1}{2k-j+1}.$$

Consequently,

$$\sum_{i=2}^n ((n+i)^{2k} - i^{2k}) < 2^{2k} (n+1)^{2k+1} < 8^{2k} n^{2k+1}.$$

Since  $\sigma^2 = \frac{n^3-n}{6} > \frac{n^3}{8}$  when  $n$  is sufficiently large, we have

$$\sigma^{-2k} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}) < 64^{2k} n^{1-k} \leq n^{-1/3} 64^{2k} n^{-k/3},$$

for large  $n$  and  $k > 1$ . Thus

$$\begin{aligned} & \left| \sum_{2 \nmid k, k \geq 3} B_{2k} \frac{t^{2k}}{2k(2k)! \sigma^{2k}} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}) \right| \\ & < n^{-1/3} \sum_{2 \nmid k, k \geq 3} |B_{2k}| \frac{t^{2k}}{2k(2k)!} 64^{2k} n^{-k/3} \\ & = n^{-1/3} \sum_{2 \nmid k, k \geq 3} |B_{2k}| \frac{(64tn^{-\frac{1}{6}})^{2k}}{2k(2k)!}. \end{aligned}$$

In view of the following asymptotic expansion of the Bernoulli numbers [1],

$$|B_{2n}| \sim \frac{2(2n)!}{(2\pi)^{2n}},$$

the convergent radius  $R$  of the series  $\sum_{2 \nmid k, k \geq 3} |B_{2k}| \frac{t^{2k}}{2k(2k)!}$  equals  $2\pi$ . Since  $t$  is from a bounded set, when  $n$  is large enough, the series

$$\sum_{2 \nmid k, k \geq 3} |B_{2k}| \frac{(64tn^{-\frac{1}{6}})^{2k}}{2k(2k)!}$$

converges. Moreover, it is evident that  $64tn^{-\frac{1}{6}} < 1$ , we can bound the above summation by the constant

$$M_1 = \sum_{2 \nmid k, k \geq 3} |B_{2k}| \frac{1}{2k(2k)!}.$$

Similarly, it can be deduced that

$$\sum_{2 \nmid k, k \geq 2} B_{2k} \frac{t^{2k}}{2k(2k)! \sigma^{2k}} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}) < \frac{M_2}{n^{\frac{1}{3}}},$$

where  $M_2 = \sum_{2|k, k \geq 2} B_{2k} \frac{1}{2k(2k)!}$  is a constant. Hence

$$\sum_{k=2}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)! \sigma^{2k}} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}) < \frac{M_1 + M_2}{n^{1/3}},$$

which tends to zero as  $n \rightarrow \infty$ . This completes the proof.  $\blacksquare$

In [7], Margolius applied Bernoulli numbers to show that the distribution of the number of inversions in a random permutation is asymptotically normal. In [6], Louchard and Prodinger used the saddle point method to derive some stronger results. Based on Lemma 2.2, we obtain the following theorem.

**Theorem 2.3.** *When  $n \rightarrow \infty$ , the random variable*

$$\eta_n = \frac{\xi_n - E(\xi_n)}{\text{Var}(\xi_n)^{\frac{1}{2}}}$$

*has the standard normal distribution.*

*Proof.* Let  $M_n(q)$  denote the moment generating function of  $\xi_n$ . Then we have  $M_n(q) = \phi_n(e^q)$ , see Sachkov [8]. Hence

$$\begin{aligned} M_n(q) &= \frac{n+1}{\binom{2n}{n}} \frac{1-e^q}{1-e^{(n+1)q}} \cdot \prod_{i=1}^n \frac{1-e^{(n+i)q}}{1-e^{iq}} \\ &= \prod_{i=2}^n \frac{i}{n+i} \cdot \prod_{i=2}^n \frac{1-e^{(n+i)q}}{1-e^{iq}} \\ &= \prod_{i=2}^n \frac{(1-e^{(n+i)q})/(n+i)}{(1-e^{iq})/i} \\ &= \exp \left\{ \frac{1}{2} \sum_{i=2}^n ((n+i)q - iq) \right\} \prod_{i=2}^n \frac{(e^{(n+i)q/2} - e^{-(n+i)q/2}) / \frac{n+i}{2}}{(e^{iq/2} - e^{-iq/2}) / \frac{i}{2}} \\ &= \exp \left\{ \frac{n(n-1)q}{2} \right\} \prod_{i=2}^n \frac{\sinh((n+i)q/2) / \frac{n+i}{2}}{\sinh(iq/2) / \frac{i}{2}}. \end{aligned}$$

Recalling the following relation on the Bernoulli numbers [7]

$$(2.2) \quad \ln \left( \frac{\sinh(x/2)}{x/2} \right) = \sum_{k=1}^{\infty} B_{2k} \frac{x^{2k}}{2k(2k)!},$$

we find that

$$\begin{aligned} \ln M_n(q) &= \frac{n(n-1)}{2} q + \sum_{i=2}^n \left( \ln \left( \frac{\sinh((n+i)q/2)}{(n+i)/2} \right) - \ln \left( \frac{\sinh(iq/2)}{i/2} \right) \right) \\ &= \frac{n(n-1)}{2} q + \sum_{k=1}^{\infty} B_{2k} \frac{q^{2k}}{2k(2k)!} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}). \end{aligned}$$

Setting  $q = t/\sigma$ , where  $\sigma$  is the standard deviation of  $\xi_n$  as given in Theorem 2.1, we are led to the expansion

$$\ln M_n(t/\sigma) = \frac{n(n-1)t}{2\sigma} + \sum_{k=1}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)!\sigma^{2k}} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}).$$

Applying Lemma 2.2, we have, when  $n \rightarrow \infty$ ,

$$\sum_{k=2}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)!\sigma^{2k}} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}) \rightarrow 0$$

uniformly for  $t$  from any bounded set. Finally,

$$\begin{aligned} & \lim_{n \rightarrow \infty} M_n(t/\sigma) \exp \left\{ -\frac{n(n-1)t}{2\sigma} \right\} \\ &= \lim_{n \rightarrow \infty} \exp \left\{ \sum_{k=1}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)!\sigma^{2k}} \sum_{i=2}^n ((n+i)^{2k} - i^{2k}) \right\} \\ &= \lim_{n \rightarrow \infty} \exp \left\{ B_2 \frac{t^2}{2(2)!\sigma^2} \sum_{i=2}^n ((n+i)^2 - i^2) \right\} \\ &= e^{t^2/2}, \end{aligned}$$

which coincides with the moment generating function of the standard normal distribution. Employing Curtiss's theorem [8], we reach the conclusion that  $\eta_n$  has the standard normal distribution when  $n$  approaches infinity. ■

### 3. A GENERAL SETTING

In this section, we will determine the limiting distribution of the coefficients of a quotient of products and will give two special cases.

**Theorem 3.1.** *Let  $a_1, a_2, a_3, \dots$  and  $b_1, b_2, b_3, \dots$  be two sequences of positive numbers, and let*

$$\phi_n(x) = \sum_k p_n(k) x^k = \frac{(1 - q^{a_1})(1 - q^{a_2}) \cdots (1 - q^{a_n})}{(1 - q^{b_1})(1 - q^{b_2}) \cdots (1 - q^{b_n})}.$$

*Suppose that  $\xi_n$  is the random variable corresponding to the generating function  $\phi_n(x)$ , that is,*

$$P(\xi_n = k) = \frac{p_n(k)}{\sum_k p_n(k)}.$$

*Then  $\xi_n$  is normally distributed as  $n \rightarrow \infty$ , if and only if*

$$\sum_{k=1}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)!} \left( \sum_{i=1}^n (a_i^{2k} - b_i^{2k}) \right) \frac{1}{\left( \sum_{i=1}^n (a_i^2 - b_i^2) \right)^k} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

*Proof.* The expectation of  $\xi_n$  is easy to compute, as given below:

$$E(\xi_n) = \phi_n(x)'_{q=1} = \frac{1}{2} \sum_{i=1}^n (a_i - b_i).$$

Proceeding analogously as in the proof of Theorem 2.1, we find

$$(3.1) \quad \sigma^2 = \text{Var}(\xi_n) = \frac{1}{12} \sum_{i=1}^n (a_i^2 - b_i^2).$$

Hence,

$$B_2 \frac{t^2}{2(2)!\sigma^2} \left( \sum_{i=1}^n (a_i^2 - b_i^2) \right) = \frac{1}{6} \cdot \frac{t^2}{4 \cdot \frac{1}{12} \left( \sum_{i=1}^n (a_i^2 - b_i^2) \right)} \cdot \left( \sum_{i=1}^n (a_i^2 - b_i^2) \right) = \frac{t^2}{2}.$$

By the same procedure as in the proof of Theorem 2.3, we obtain

$$\begin{aligned} & \lim_{n \rightarrow \infty} M_n(t/\sigma) \exp \left\{ \frac{1}{2} \sum_{i=1}^n (a_i^{2k} - b_i^{2k}) \right\} \\ &= e^{t^2/2} \lim_{n \rightarrow \infty} \exp \left\{ \sum_{k=2}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)!\sigma^{2k}} \left( \sum_{i=1}^n (a_i^{2k} - b_i^{2k}) \right) \right\}. \end{aligned}$$

It follows that the limiting distribution of  $p_n(k)$  is normal if and only if

$$(3.2) \quad \sum_{k=2}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)!\sigma^{2k}} \left( \sum_{i=1}^n (a_i^{2k} - b_i^{2k}) \right) \rightarrow 0 \quad \text{as } n \rightarrow \infty,$$

for  $t$  from any bounded set. By virtue of the variance formula (3.1), the condition (3.2) is equivalent to

$$(3.3) \quad \sum_{k=1}^{\infty} B_{2k} \frac{t^{2k}}{2k(2k)!} \frac{\sum_{i=1}^n (a_i^{2k} - b_i^{2k})}{\left( \sum_{i=1}^n (a_i^2 - b_i^2) \right)^k} \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

for  $t$  from any bounded set. Thus (3.2) is verified. This completes the proof.  $\blacksquare$

**Corollary 3.2.** *Let  $p_n(k)$  be given as in the above theorem. Suppose that for  $k \geq 2$ , there exist constants  $\alpha > 0$ ,  $\beta < 0$  and  $\gamma < 0$  such that*

$$(3.4) \quad \frac{\sum_{i=1}^n (a_i^{2k} - b_i^{2k})}{\left( \sum_{i=1}^n (a_i^2 - b_i^2) \right)^k} < n^\gamma (\alpha n^\beta)^{2k},$$

for  $t$  from any bounded set. Then the limiting distribution of  $p_n(k)$  is normal.

*Proof.* Note that the convergent radius  $R$  of the series

$$\sum_{2 \nmid k, k \geq 3} |B_{2k}| \frac{x^{2k}}{2k(2k)!}$$

is  $2\pi$ . If (3.4) holds for  $k > 1$ , then for  $t$  from any bounded set, and for sufficiently large  $n$ , we have

$$\left| t^{2k} \sum_{i=1}^n (a_i^{2k} - b_i^{2k}) / \sigma^{2k} \right| \leq n^\gamma (t\alpha n^\beta)^{2k},$$

where  $t\alpha n^\beta < 2\pi$ . It is clear that  $n^\gamma \rightarrow 0$  since  $\gamma < 0$ .  $\blacksquare$

If we choose  $\alpha = 32\sqrt{3}/3$ ,  $2\beta = \gamma = -\frac{1}{3}$ , Theorem 3.2 contains Theorem 2.3 as a special case. We now give two more examples. One is the following  $q$ -analog of the Catalan numbers

$$c_n(q) = \frac{[2]}{[2n]} \begin{bmatrix} 2n \\ n-1 \end{bmatrix},$$

which are symmetric and unimodal, see Stanley [9].

Using Theorem 3.1, we reach the following assertion.

**Corollary 3.3.** *The distribution of the coefficients in  $c_n(q)$  is asymptotically normal.*

*Proof.* First, we write  $c_n(q)$  in the following form:

$$\frac{\prod_{i=3}^n (1 - q^{n+i-1})}{(1-q) \prod_{i=3}^{n-1} (1 - q^i)},$$

Set  $a_1 = a_2 = 1$ ,  $a_i = n + i - 1$ ,  $3 \leq i \leq n$ , and  $b_1 = b_2 = 1$ ,  $b_3 = 1$ ,  $b_i = i - 1$ ,  $4 \leq i \leq n$ . Then we have

$$\begin{aligned} \sum_{i=1}^n (a_i^{2k} - b_i^{2k}) &= (a_3^{2k} - b_3^{2k}) + \sum_{i=4}^n (a_i^{2k} - b_i^{2k}) \\ &= (n+2)^{2k} - 1 + \sum_{i=3}^{n-1} ((n+i)^{2k} - i^{2k}) \end{aligned}$$

and

$$\begin{aligned} \left( \sum_{i=1}^n (a_i^2 - b_i^2) \right)^k &= \left( (n+2)^2 - 1 + \sum_{i=3}^{n-1} ((n+i)^2 - i^2) \right)^k \\ &= (n-1)^k (n+1)^k (2n-3)^k. \end{aligned}$$

By the same arguments as in the proof of Lemma 2.2, we may set  $\alpha = 32\sqrt{3}/3$  and  $2\beta = \gamma = -\frac{1}{3}$  such that the condition (3.4) is satisfied. Therefore, Theorem 3.1 implies the limiting distribution of the coefficients of  $c_n(q)$ .  $\blacksquare$

The  $m$ -Catalan numbers are defined by

$$C_{n,m} = \frac{1}{(m-1)n+1} \binom{mn}{n},$$

for  $n \geq 1$ . Accordingly, the generalized  $q$ -Catalan numbers are given by

$$C_{n,m}(q) = \frac{1}{[(m-1)n+1]} \begin{bmatrix} mn \\ n \end{bmatrix}.$$

Theorem 3.1 has the following consequence.

**Corollary 3.4.** *The coefficients of the generalized  $q$ -Catalan numbers  $C_{n,m}(q)$  are normally distributed when  $n \rightarrow \infty$ .*

*Proof.* First, express  $C_{n,m}(q)$  as follows

$$\prod_{i=2}^n \frac{1 - q^{(m-1)n+i}}{1 - q^i}.$$

Set  $a_1 = 1$ ,  $a_i = (m-1)n + i$ ,  $2 \leq i \leq n$ , and  $b_1 = 1$ ,  $b_i = i$ ,  $2 \leq i \leq n$ . Then we have

$$\sum_{i=1}^n (a_i^{2k} - b_i^{2k}) = \sum_{i=2}^n (a_i^{2k} - b_i^{2k}) = \sum_{i=2}^n \sum_{j=1}^{2k} \binom{2k}{j} ((m-1)n)^{2k-j} i^j.$$

The same argument as in the proof of Lemma 2.2 yields the following bound

$$\sum_{i=1}^n (a_i^{2k} - b_i^{2k}) < 8^{2k} ((m-1)n)^{2k+1}.$$

Now,

$$\begin{aligned} \left( \sum_{i=1}^n (a_i^2 - b_i^2) \right)^k &= \left( \sum_{i=2}^n (((m-1)n + i)^2 - i^2) \right)^k \\ &> (m-1)^{2k} n^{2k} (n-1)^k \\ &> (m-1)^{2k+1} n^{3k} / (2m)^k. \end{aligned}$$

It follows that

$$\frac{\sum_{i=1}^n (a_i^{2k} - b_i^{2k})}{\left( \sum_{i=1}^n (a_i^2 - b_i^2) \right)^k} < (8\sqrt{2m})^{2k} n^{1-k}.$$

Again, by the same arguments as in the proof of Lemma 2.2, we may set  $\alpha = 8\sqrt{2m}$  and  $2\beta = \gamma = -\frac{1}{3}$  such that the condition (3.4) holds. Finally, we may use Theorem 3.1 to get the desired distribution.  $\blacksquare$

#### 4. OPEN PROBLEMS

While the  $q$ -Catalan numbers are not unimodal for small  $n$ , see Stanley [9], the limiting distribution suggests that the coefficients are almost unimodal in certain sense for sufficiently large  $n$ . Obviously, the first and the last term should not be taken into account; otherwise one can never expect to have unimodality. In fact, an easy computation indicates that  $C_n(q)$  are unimodal for  $n \geq 16$ .

**Conjecture 4.1.** *The sequence  $\{m_n(1), \dots, m_n(n(n-1)-1)\}$  is unimodal when  $n$  is sufficiently large.*

When  $n > 70$ , numerical evidence is suggestive of a stronger conjecture:

**Conjecture 4.2.** *There exists an integer  $t$  such that when  $n$  is sufficiently large, the sequence  $\{m_n(t), \dots, m_n(n(n-1)-t)\}$  is log-concave, namely,*

$$(m_n(k))^2 \geq m_n(k+1)m_n(k-1)$$

for  $t+1 \leq k \leq n(n-1)-t-1$ . Moreover, the minimum value of  $t$  seems to be 75.

We would also conjecture that similar properties hold for the generalized  $q$ -Catalan numbers.

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